

# Statistical Model Checking and Rare Events

**Paolo Zuliani**

Joint work with Edmund M. Clarke  
Computer Science Department, CMU

# Probabilistic Verification

- Verification of **stochastic system models** via **statistical model checking**
- **Temporal logic** specification:
  - “the amount of p53 exceeds  $10^5$  within 20 minutes”
- If  $\phi =$  “p53 exceeds  $10^5$  within 20 minutes”

Probability ( $\phi$ ) = ?

# Equivalently

- A biased coin (Bernoulli random variable):
  - Prob (Heads) =  $p$                       Prob (Tails) =  $1-p$
  - $p$  is unknown
- Question: What is  $p$ ?
- A solution: flip the coin a number of times, collect the outcomes, and use statistical estimation

# Statistical Model Checking

## Key idea

(Haakan Younes, 2001)

- System behavior w.r.t. **property**  $\phi$  can be modeled by a **Bernoulli random variable** of parameter  $p$ :
  - System satisfies  $\phi$  with (**unknown**) probability  $p$
- Question: What is  $p$ ?
- Draw a sample of system **simulations** and use:
  - Statistical **estimation**: returns “ $p$  in interval  $(a,b)$ ” with high probability

# Statistical Model Checking

- Statistical Model Checking is a **Monte Carlo** method
- Problems arise when  $p$  is **very small** (rare event)
- The **number of simulations** (coin flips) needed to estimate  $p$  accurately grows **too large**
- Need to deal with this ...

# Rare events

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- **Standard (Crude) Monte Carlo**: generate  $K$  i.i.d. samples of  $X$ ; return the estimator  $e_K$

$$e_K = \frac{1}{K} \sum_{i=1}^K I(X_i \geq t) = \frac{k_t}{K}$$

- $\text{Prob}(e_K \rightarrow p_t) = 1$  for  $K \rightarrow \infty$  (strong law LN)

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- $\text{Var}[e_K] = \frac{p_t(1-p_t)}{K}$
- By the Central Limit Theorem (CLT), the **distribution of  $e_K$**  converges to a normal distribution with:
  - mean  $p_t$
  - variance  $\frac{p_t(1-p_t)}{K}$
- **Relative Error (RE)** =  $\frac{\sqrt{\text{var}[e_K]}}{E[e_K]} = \frac{\sqrt{p_t(1-p_t)}}{p_t\sqrt{K}}$

# Rare events

- $RE = \frac{\sqrt{p_t(1-p_t)}}{p_t\sqrt{K}}$
- Fix  $K$ , then RE is unbounded as  $p_t \rightarrow 0$
- More accuracy  $\rightarrow$  more samples
- Want confidence interval of relative accuracy  $\delta$  and coverage probability  $c$ , i.e., estimate  $e_K$  must satisfy:

$$\text{Prob}(|e_K - p_t| < \delta \cdot p_t) \geq c$$

- How many samples do we need?

# Rare events

- From the CLT, a 99% (approximate) confidence interval of **relative accuracy**  $\delta$  needs about

$$K \approx \frac{1 - p_t}{p_t \delta^2} \text{ samples}$$

Thus,  $\text{Prob}(|e_K - p_t| < \delta p_t) \approx 0.99$

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- Examples:
  - $p_t = 10^{-9}$  and  $\delta = 10^{-2}$  (ie, 1% relative accuracy) we need about  $10^{13}$  samples!!
  - Bayesian estimation requires about  $6 \times 10^6$  samples with  $p_t = 10^{-4}$  and  $\delta = 10^{-1}$

# A solution

- Importance Sampling (1940s)
- A variance-reduction technique
- Can result in dramatic reduction in sample size

# Importance Sampling

- The fundamental Importance Sampling identity

$$\begin{aligned} p_t &= E[I(X \geq t)] \\ &= \int I(x \geq t) f(x) dx \\ &= \int I(x \geq t) \frac{f(x)}{f_*(x)} f_*(x) dx \\ &= \int I(x \geq t) W(x) f_*(x) dx \\ &= E_*[I(X \geq t) W(X)] \end{aligned}$$

$f$  is the density of  $X$

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likelihood ratio

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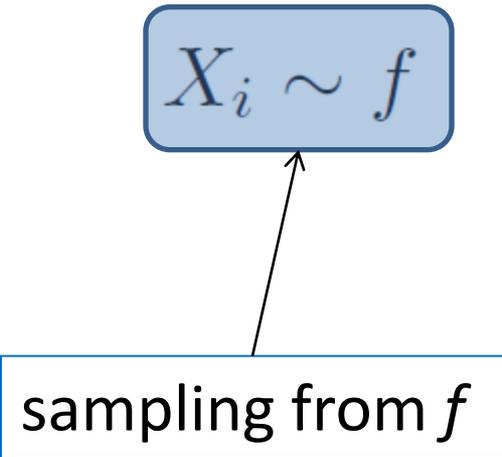
- Estimate  $p_t = E[X \geq t] = \text{Prob}(X \geq t)$
- A sample  $X_1, \dots, X_K$  iid as  $f$
- The **crude Monte Carlo** estimator is

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A diagram consisting of two rectangular boxes. The top box is light blue with rounded corners and contains the text  $X_i \sim f$ . The bottom box is white with a thin blue border and contains the text "sampling from  $f$ ". A black arrow points from the top of the bottom box to the bottom of the top box.

$$X_i \sim f$$

sampling from  $f$

# Importance Sampling

- Define a **biasing density**  $f_*$
- Compute the **IS estimator**

$$\hat{p}_t = \frac{1}{K} \sum_{i=1}^K I(X_i \geq t) W(X_i), \quad X_i \sim f_*$$

where  $W(x) = \frac{f(x)}{f_*(x)}$  is the **likelihood ratio**

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# Importance Sampling

- Need to choose a “good” **biasing density** (low variance)

- **Optimal density:**  $f_*(x) = \frac{I(x \geq t)f(x)}{p_t}$

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# Cross-Entropy Method

(R. Rubinstein)

- Suppose the density of  $X$  in a family of densities  $\{f(\cdot; \nu)\}$ 
  - the “nominal”  $f$  is  $f(x; u)$
- **Key idea**: choose a parameter  $\nu$  such that the *distance* between  $f_*$  and  $f(\cdot; \nu)$  is minimal
- The Kullback-Leibler divergence (cross-entropy) is a measure of “distance” between two densities
- First used for rare event simulation by Rubinstein (1997)

# Cross-Entropy Method

- The **KL divergence (cross-entropy)** of densities  $g, h$  is

$$D(g, h) = E_g \left[ \ln \frac{g(X)}{h(X)} \right] = \int g(x) \ln g(x) dx - \int g(x) \ln h(x) dx$$

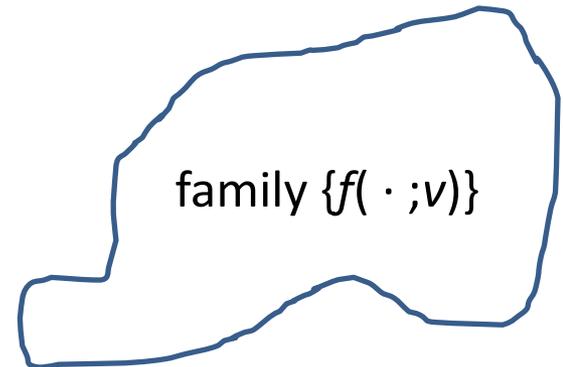
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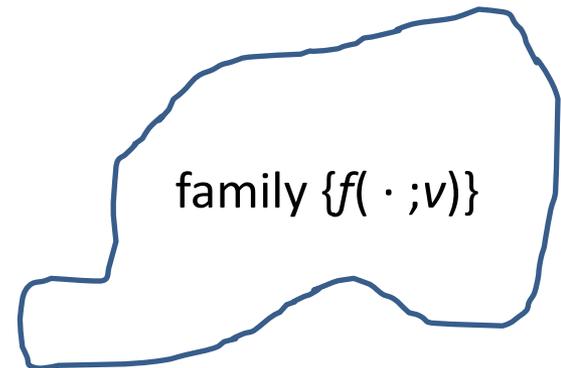
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optimal density  $f_*$

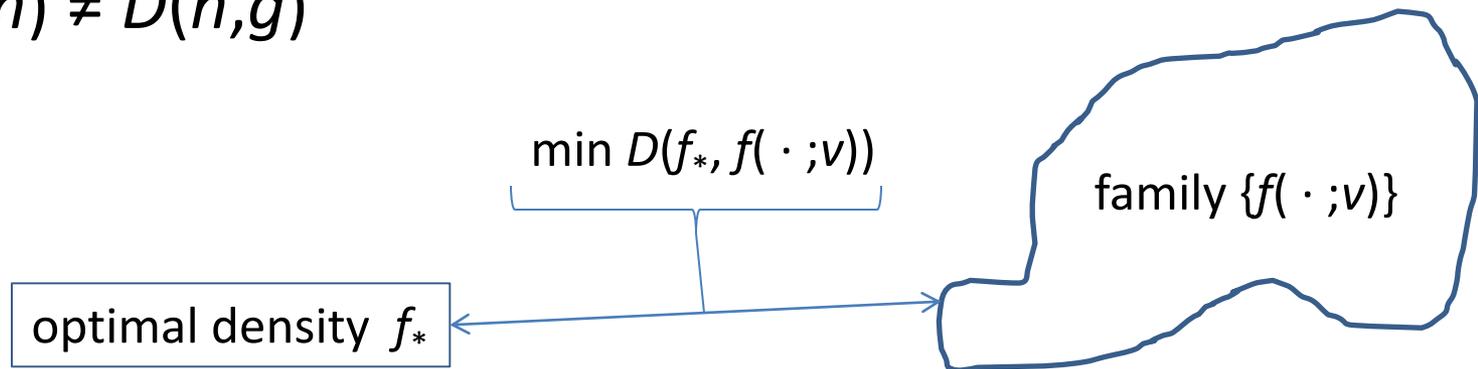


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- Step 2 is “easy”
- Step 1 is **not so easy**

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$$V_* =$$

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$$\mathbf{v}_* = \arg \min_{\mathbf{v}} E_{f_*} \left[ \ln \frac{f_*(X)}{f(X; \mathbf{v})} \right] = \arg \min_{\mathbf{v}} \int f_*(x) \ln f_*(x) dx - \int f_*(x) \ln f(x; \mathbf{v}) dx$$

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# Cross-Entropy Method

- For certain families  $\{f(\cdot; \nu)\}$  (eg, one-dim exponential) the problem

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can be solved **analytically**:

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- **In practice:** get  $X_1, \dots, X_K$  samples iid as  $f(\cdot; u)$  and compute the approximation

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In general, one would have to (numerically) solve the problem

$$\frac{1}{K} \sum_{i=1}^K I(X_i \geq t) \nabla \ln f(X_i; v) = 0$$

# Cross-Entropy Method

- **Problem:** If  $\{X \geq t\}$  is a rare event, then this fails

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- Most terms in both sums will be zero!

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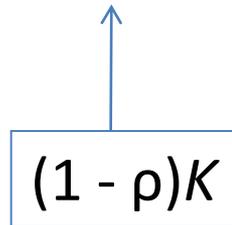
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- Now compute  $v_*$  “as usual”. Iterate until  $t' = t$

# Cross-Entropy with Rare Events

- Does NOT work with statistical model checking
- **Problem:** sample quantile computation
- Order the sample performances

$$S_{(1)} \leq \dots \leq S_{(i)} \leq \dots \leq S_{(K)}$$



$(1 - \rho)K$

- In statistical model checking, sample performances are either 0 (property false) or 1 (property true)

# Cross-Entropy with Rare Events

- However ...

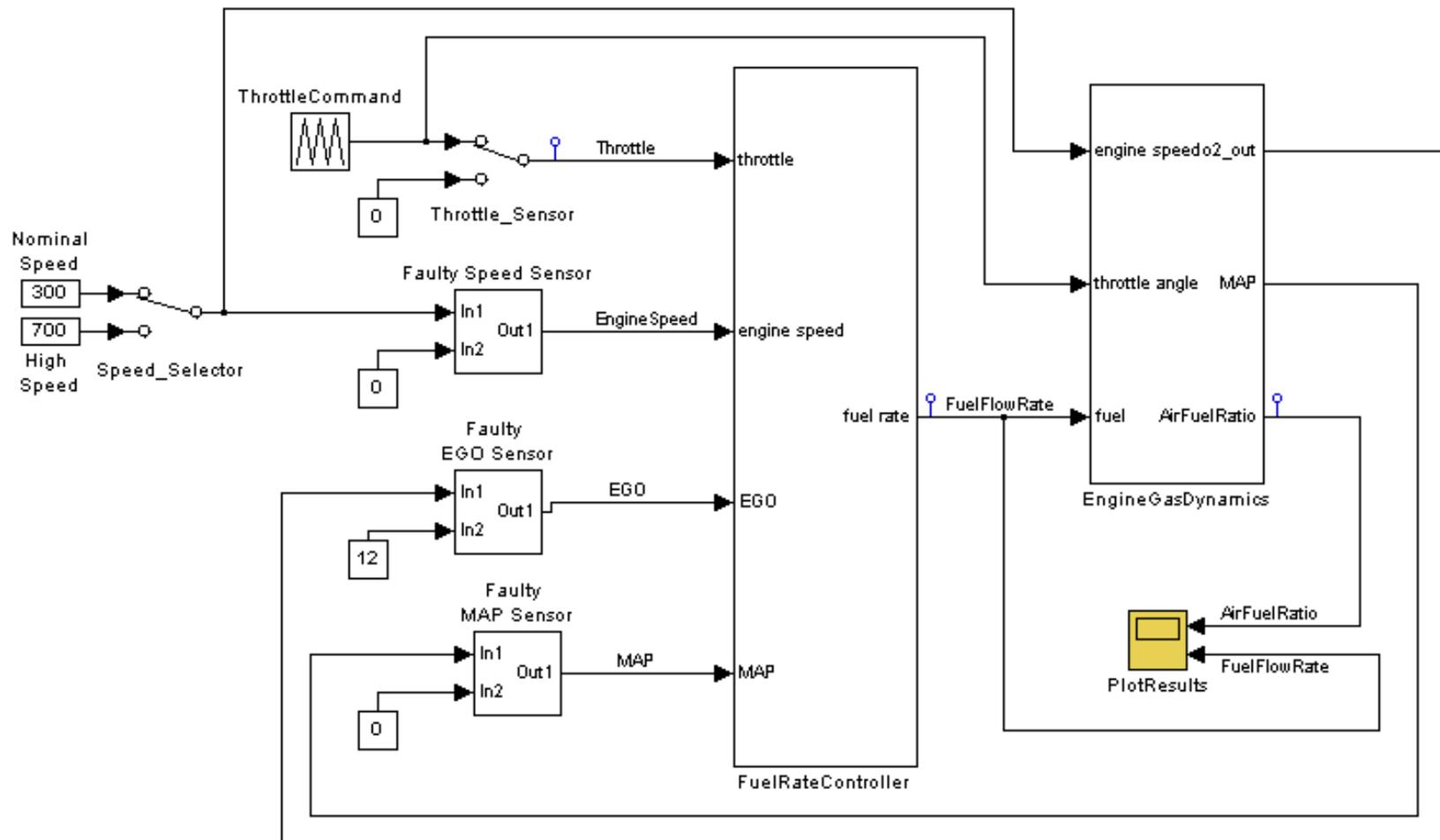
$$v_* = \frac{E_u[I(X \geq t)X]}{E_u[I(X \geq t)]} = \frac{E_w[I(X \geq t)W(X;u,w)X]}{E_w[I(X \geq t)W(X;u,w)]}$$

where  $W(x;u,w) = \frac{f(x;u)}{f(x;w)}$  for an arbitrary parameter  $w$

Work in progress

# Example: Fuel Control System

## The Stateflow/Simulink model



# Verification

- We want to estimate the probability that

$$\mathcal{M}, \text{FaultRate} \models F^{100} G^1(\text{FuelFlowRate} = 0)$$

- “It is the case that within 100 seconds, FuelFlowRate is zero for 1 second”
- $\text{FaultRate} = 1/3600s$  (same value for the three sensors)

# Importance Sampling

- Ran cross-entropy method to estimate optimal biasing density with  $FaultRates = \{1/7, 1/8, 1/9\}$
- Used 100 samples for this, and obtained
  - $NewRates_* = \{1/2.007, 1/1.0113, 1/1.7277\}$
- Run importance sampling with 1,000 samples and  $NewRates_*$ 
  - Probability estimate  $9.1855 \times 10^{-15}$

# Conclusions

- Need to be able to deal with rare events in statistical model checking
- The Cross-Entropy method is an interesting, semi-automatic technique
- Research: adaptive technique for stat. model checking
- [Further benefit: cross-entropy method also applies to optimization, *eg*, finding policies for MDPs]

**The End**

Questions?